

Cross-validation in Scattered Data Approximation

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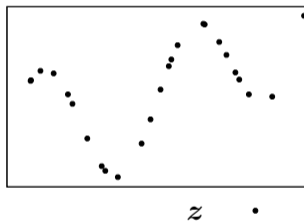
4th of May 2021



Mathematik!
TU Chemnitz

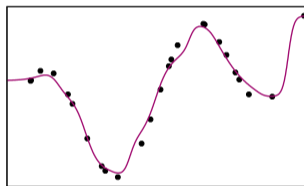
Setting

- data $\mathbf{z} = (\mathbf{x}_i, y_i)_{i=1}^n \in (\Omega \times Y)^n$
- (\mathbf{x}_i, y_i) distributed according to ρ



Setting

- data $z = (\mathbf{x}_i, y_i)_{i=1}^n \in (\Omega \times Y)^n$
- (\mathbf{x}_i, y_i) distributed according to ρ
- reconstruction algorithm
 $R_h: (\Omega \times Y)^n \rightarrow Y^\Omega$



$R_h(z)$ —•—

Problem 1: Estimate Risk

Given: data \mathbf{z} , reconstruction model $R_h(\mathbf{z})$

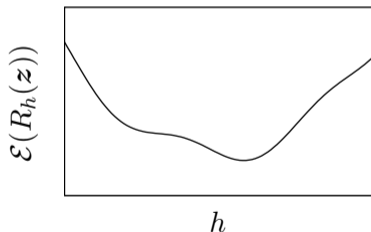
Question: How big is the L_2 -reconstruction error of $R_h(\mathbf{z})$?

$$\mathcal{E}(R_h(\mathbf{z})) = \int_{\Omega \times Y} |(R_h(\mathbf{z}))(\mathbf{x}) - y|^2 \, d\rho(\mathbf{x}, y)$$

Problem 2: Model selection

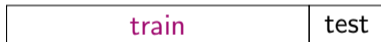
Given: data z , hypothesis space $\{R_h(z) : h \in H\}$

Question: Which reconstruction model $R_h(z)$ to choose?



Purly data-driven method to approximate the risk

- validation set



Purly data-driven method to approximate the risk

- validation set
- "Probably the simplest and most widely used method for estimation prediction error is cross-validation." [Hastie 2001]

test	train		
train	test	train	
train		test	train
train			test
train			test

Leave-one-out cross-validation [Stone 1974]

- 1 calculate the reconstructions $R_h(\mathbf{z}_{-i})$ from the function values $\mathbf{z}_{-i} = (z_1, \dots, z_{i-1}, z_{i+1}, \dots, z_n)$
- 2 evaluate the residual of $R_h(\mathbf{z}_{-i})$ in the i -th node \mathbf{x}_i

$$|R_h(\mathbf{z}_{-i})(\mathbf{x}_i) - y_i|$$

- 3 calculate the mean value with respect to all nodes

$$\text{CV}(\mathbf{z}, \lambda) := \frac{1}{n} \sum_{i=1}^n |R_h(\mathbf{z}_{-i})(\mathbf{x}_i) - y_i|^2$$

Table of contents

- ① Fast computation
- ② Theoretical foundation

- **Deshpande and Girard.** Fast computation of cross-validated robust splines and other non-linear smoothing splines. 2010.
- **Mullin and Sukthankar.** Complete Cross-Validation for Nearest Neighbor Classifiers. 2000.
- **An, Liu, and Venkatesh.** Fast cross-validation algorithms for least squares support vector machine and kernel ridge regression. 2007.
- ...
- cross-validation has 4 000 000 hits on Google Scholar
- fast cross-validation has 3 000 000 hits on Google Scholar

Penalized least squares estimation

$$R_h(\mathbf{z}) = \sum_{\mathbf{k} \in I} \hat{g}(\mathbf{k}) \varphi_{\mathbf{k}}(\cdot)$$

where

$$\begin{aligned} \hat{\mathbf{g}} &= \arg \min_{\hat{\mathbf{g}}} \|\mathbf{F}\hat{\mathbf{g}} - \mathbf{y}\|_{\mathbf{W}}^2 + \lambda \|\hat{\mathbf{g}}\|_{\hat{\mathbf{W}}}^2 \\ &= \left(\mathbf{F}^* \mathbf{W} \mathbf{F} + \lambda \hat{\mathbf{W}} \right)^{-1} \mathbf{F}^* \mathbf{W} \mathbf{y}. \end{aligned}$$

for

- $\mathbf{F} = (\varphi_{\mathbf{k}}(\mathbf{x}_i))_{i=1, \dots, n, \mathbf{k} \in I}$
- $\mathbf{y} = (y_i)_{i=1}^n$
- $\mathbf{W} = \text{diag}(w_i)_{i=1}^n$
- $\hat{\mathbf{W}} = \text{diag}(\hat{w}(\mathbf{k}))_{\mathbf{k} \in I}$

Evaluating the reconstruction for full data, we obtain

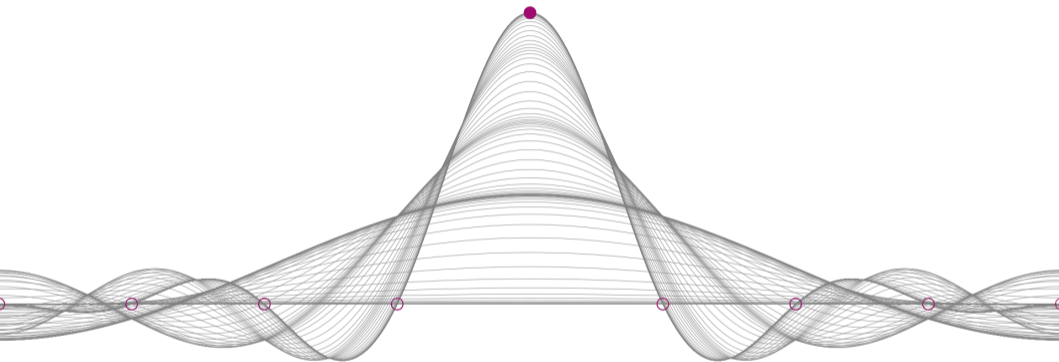
$$\begin{aligned}\mathbf{F}\hat{\mathbf{g}}_\lambda &= \mathbf{F} \left(\mathbf{F}^* \mathbf{W} \mathbf{F} + \lambda \hat{\mathbf{W}} \right)^{-1} \mathbf{F}^* \mathbf{W} \mathbf{y} \\ &= \mathbf{H}_\lambda \mathbf{y}.\end{aligned}$$

Theorem

[Golub, Heath, Whaba 1979]

$$\text{CV}(\mathbf{z}, \lambda) = \frac{1}{n} \sum_{i=1}^n \left| \frac{[\mathbf{H}_\lambda \mathbf{y} - \mathbf{y}]_i}{1 - h_{i,i}} \right|^2$$

with $h_{i,i}$ being the diagonal entries of \mathbf{H}_λ .



$$h_{i,i} = e_i^T \mathbf{H}_\lambda e_i$$

For $\Omega = \mathbb{T}^d$, exponential functions

$$\varphi_{\mathbf{k}} = \exp(2\pi i \mathbf{k} \cdot \mathbf{x}),$$

the full grid

$$\mathbf{X} = \left\{ \frac{\mathbf{m}}{N} : \mathbf{m} \in \mathbb{Z}^d \cap \prod_{t=1}^d [0, N) \right\}$$

the diagonal entries can be computed via

$$h_{i,i} = \frac{1}{N^d} \sum_{\mathbf{k} \in I} \frac{1}{1 + \lambda \hat{w}(\mathbf{k})}.$$

Theorem (Diagonal entries $h_{i,i}$ of the hat matrix) [B, Hielscher, Potts, 2019]

Let I be an index set and \mathbf{X} , \mathbf{W} form an exact quadrature rule for $\exp(2\pi i \mathbf{k} \cdot \mathbf{x})$ with $\mathbf{k} \in \mathcal{D}(I) = \{\mathbf{k}_1 - \mathbf{k}_2 : \mathbf{k}_1, \mathbf{k}_2 \in I\}$. Then

$$h_{i,i} = w_i \sum_{\mathbf{k} \in I} \frac{1}{1 + \lambda \hat{w}(\mathbf{k})}.$$

Theorem (Diagonal entries $h_{i,i}$ of the hat matrix) [B, Hielscher, Potts, 2019]

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$$h_{i,i} = w_i \sum_{\mathbf{k} \in I} \frac{1}{1 + \lambda \hat{w}(\mathbf{k})}.$$

Proof:

Using quadrature, we have $\mathbf{F}^* \mathbf{W} \mathbf{F} = \mathbf{I}$ and thus

$$\begin{aligned} h_{i,i} &= \left(\mathbf{F} \left(\mathbf{F}^* \mathbf{W} \mathbf{F} + \lambda \hat{\mathbf{W}} \right)^{-1} \mathbf{F}^* \mathbf{W} \right)_{i,i} \\ &= \left(\mathbf{F} \operatorname{diag} \left(\frac{1}{1 + \lambda \hat{w}(\mathbf{k})} \right)_{\mathbf{k} \in I} \mathbf{F}^* \mathbf{W} \right)_{i,i} = w_i \sum_{\mathbf{k} \in I} \frac{1}{1 + \lambda \hat{w}(\mathbf{k})}. \end{aligned}$$

Algorithm: Computation of the cross-validation score

Input:

- \mathbf{X}
- $\mathbf{W} = \text{diag}(w_i)_{i=1}^n$
- $\hat{\mathbf{W}}$
- $\mathbf{y} = (y_i)_{i=1}^n$
- λ

Output:

- $\text{CV}(\mathbf{z}, \lambda)$

① Compute $\hat{\mathbf{g}} := \left(\mathbf{F}^* \mathbf{W} \mathbf{F} + \lambda \hat{\mathbf{W}} \right)^{-1} \mathbf{F}^* \mathbf{W} \mathbf{y}$

② Compute $\mathbf{g} := \mathbf{H}_\lambda \mathbf{y} = \mathbf{F} \hat{\mathbf{g}}$

③ Compute $h_{i,i} := w_i \sum_{\mathbf{k} \in I} \frac{1}{1 + \lambda \hat{w}(\mathbf{k})}$ for $1 \leq i \leq n$

④ Evaluate $\text{CV}(\mathbf{z}, \lambda) := \frac{1}{n} \sum_{i=1}^n \left| \frac{g_i - y_i}{1 - h_{i,i}} \right|^2$

Fast computation

domain Ω	basis $\varphi_{\mathbf{k}}$	diagonal elements $h_{i,i}$
\mathbb{T}^d	$\exp(2\pi i \mathbf{k} \cdot \mathbf{x})$	$w_i \sum_{\mathbf{k} \in I} \frac{1}{1 + \lambda \hat{w}(\mathbf{k})}$
$[0, 1]$	$\cos(k \arccos x)$	$\frac{w_i}{2} \left(\frac{2}{\pi + \lambda \hat{w}_0} + \sum_{k=1}^{K-1} \frac{\cos(2k \arccos x_i) + 1}{\pi/2 + \lambda \hat{w}_k} \right)$
\mathbb{S}^2	$Y_{k,l}(\vartheta, \varphi)$	$\frac{w_i}{4\pi} \sum_{k=0}^K \frac{2k+1}{1 + \lambda \hat{w}_k}$
$\text{SO}(3)$	$D_k^{ll'}(\mathbf{R})$	$w_i \sum_{k=0}^K \frac{2k+1}{\frac{8\pi^2}{2k+1} + \lambda \hat{w}_k}$

[B., Hielscher, Potts 2019]

Fast matrix-vector multiplication

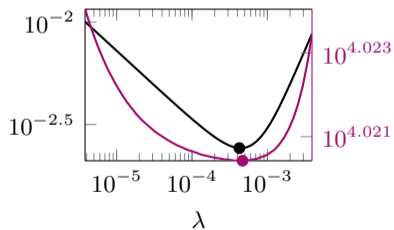
domain Ω	basis $\varphi_{\mathbf{k}}$	fast algorithm
\mathbb{T}^d	$\exp(2\pi i \mathbf{k} \cdot \mathbf{x})$	FFT, LFFT, NFFT
$[0, 1]$	$\cos(k \arccos x)$	FCT, NFCT
\mathbb{S}^2	$Y_{k,l}(\vartheta, \varphi)$	NFSFT
$\text{SO}(3)$	$D_k^{ll'}(\mathbf{R})$	NSOFT

Thus, $\text{CV}(\mathbf{z}, \lambda)$ computable in $\mathcal{O}(n \log n + |I|)$.

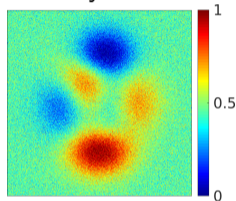
Numerical example

- 1024×1024 nodes
- 10% Gaussian noise

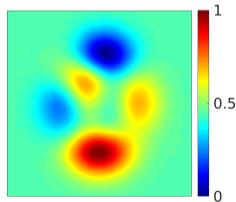
L_2 -error $\|\hat{g}_\lambda - \hat{f}\|_2$ and
cross-validation score $CV(z, \lambda)$



noisy data

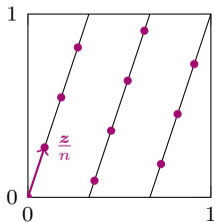


reconstruction



- $\mathbf{X} = \Lambda(\mathbf{z}, n) = \left\{ \mathbf{x} = \frac{1}{n}(m\mathbf{z} \bmod n\mathbf{1}) \in \mathbb{T}^d : m = 0, \dots, n-1 \right\}$
- Given I , there exist algorithms that find \mathbf{z} and n such that $\mathbf{F}^* \mathbf{W} \mathbf{F} = I$ for $\mathbf{W} = \text{diag}(1/n)_{i=1}^n$.
- Multiplication with \mathbf{F} can be carried out with the LFFT in $\mathcal{O}(n \log n)$.

$$\mathbf{z} = (1, 3)^T, n = 11$$



Numerical example

- $f = \bigotimes_{j=1}^7 B_2$ with 5% Gaussian noise

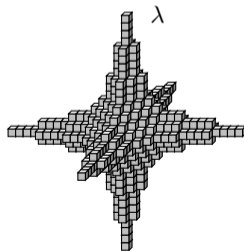
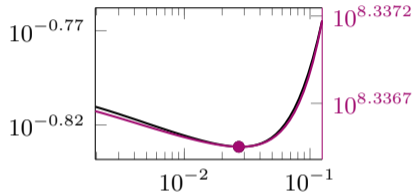
- $\hat{w}(\mathbf{k}) = \prod_{j=1}^7 \max(|n_j|^2, 1)$

- index set

$$I_{16}^{7, \text{hc}} := \left\{ \mathbf{k} \in \mathbb{Z}^7 : \prod_{j=1}^7 \max(1, |n_j|) \leq 16 \right\}$$

- \mathbf{X} : 1 105 193 rank-1 lattice nodes

L_2 -error $\|\hat{g}_\lambda - \hat{f}\|_2$ and
cross-validation score $CV(\mathbf{z}, \lambda)$



Scattered nodes

By [Tropp 2012], [Oliviera 2010], or [Rauhut 2010] we obtain for uniformly random nodes

$$F^*WF \approx I$$

with high probability, which motivates

Definition

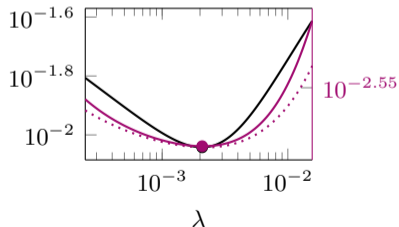
The approximated cross-validation score $\tilde{C}\tilde{V}(\lambda)$ is defined by

$$\tilde{C}\tilde{V}(\lambda) = \frac{1}{n} \sum_{i=1}^n \frac{[\mathbf{H}\mathbf{y} - \mathbf{y}]_i^2}{(1 - \tilde{h}_{i,i})^2} \quad \text{for} \quad \tilde{h}_{i,i} = w_i \sum_{\mathbf{k} \in I} \frac{1}{1 + \lambda \hat{w}(\mathbf{k})} \approx h_{i,i}.$$

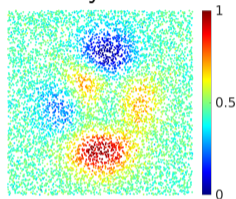
Numerical example

- 8 192 uniformly scattered nodes
- 5 % Gaussian noise

L_2 -error $\|\hat{g}_\lambda - \hat{f}\|_2$ and
cross-validation score $CV(z, \lambda)$



noisy data



reconstruction

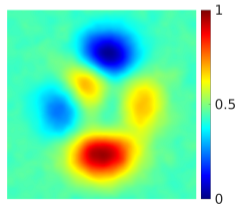


Table of contents

- ① Fast computation (for PLSE) ✓
- ② Theoretical foundation

Problem

Given: data $\mathbf{z} = (\mathbf{x}_i, f(\mathbf{x}_i))_{i=1}^n$, reconstruction model $R_h(\mathbf{z})$

Question: Relation of

$$\mathcal{E}(R_h(\mathbf{z})) = \int_{\Omega \times Y} |(R_h(\mathbf{z}))(\mathbf{x}) - f(\mathbf{x})|^2 \, d\rho(\mathbf{x}, y)$$

and

$$\text{CV}(\mathbf{z}, \lambda) = \frac{1}{n} \sum_{i=1}^n |R_h(\mathbf{z}_{-i})(\mathbf{x}_i) - y_i|^2.$$

Bound by Golub, Heath, and Whaba from 1979

For PLSE with $1/n \text{trace } \mathbf{H}_\lambda < 1$ they showed

$$\begin{aligned} & \frac{|\mathbb{E}(\mathcal{E}(R_\lambda(\mathbf{Z})) - \mathbb{E}(\text{CV}(\mathbf{Z}, \lambda)))|}{\mathbb{E}(\text{CV}(\mathbf{Z}, \lambda))} \\ & \leq \left(\frac{2}{n} \text{trace } \mathbf{H}_\lambda + \frac{(\text{trace } \mathbf{H}_\lambda)^2}{\text{trace } \mathbf{H}_\lambda^2} \right) \frac{1}{(1 - 1/n \text{trace } \mathbf{H}_\lambda)^2}. \end{aligned}$$

Bound by Li from 1987

Let

$$h^+ = \arg \min_h \mathcal{E}(R_h(\mathbf{z})) \quad \text{and} \quad h^* = \arg \min_h \text{CV}(\mathbf{z}, h).$$

Then under certain assumptions

$$\frac{\mathcal{E}(R_{h^*}(\mathbf{z}))}{\mathcal{E}(R_{h^+}(\mathbf{z}))} \rightarrow 1 \quad \text{for} \quad n \rightarrow \infty.$$

Theorem (Hoeffding 1963)

Let

- Z_1, \dots, Z_n independent rv's with values in $[0, 1]$ and
- $m = \mathbb{E} \left\{ \frac{1}{n} \sum_{i=1}^n Z_i \right\}$.

Then for $\varepsilon > 0$

$$\mathbb{P} \left\{ \left| \frac{1}{n} \sum_{i=1}^n Z_i - m \right| > \varepsilon \right\} \leq 2 \exp(-2n\varepsilon^2).$$

A function $f: \Omega^n \rightarrow \mathbb{R}$ is **c-bounded** on $\Xi \subset \Omega^n$ for $\mathbf{c} = (c_1, \dots, c_n)$, iff

$$|f(z_1, \dots, z_n) - f(z_1, \dots, z_{i-1}, z'_i, z_{i+1}, \dots, z_n)| \leq c_i$$

for all $(z_1, \dots, z_n), (z'_1, \dots, z'_n) \in \Xi$, and $1 \leq i \leq n$.

Theorem (McDiarmid 1989)

Let

- Z_1, \dots, Z_n be independent rv's with values in Ω ,
- $f: \Omega^n \rightarrow \mathbb{R}$ be c -bounded on Ω^n , and
- $m = \mathbb{E} \{f(Z_1, \dots, Z_n)\}$.

Then for $\varepsilon > 0$

$$\mathbb{P} \{|f(Z_1, \dots, Z_n) - m| > \varepsilon\} \leq 2 \exp\left(-\frac{2\varepsilon^2}{\|c\|_2^2}\right).$$

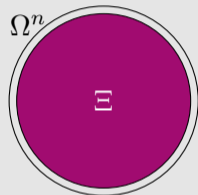
Theorem (Extension of McDiarmid by Combes 2015)

Let

- Z_1, \dots, Z_n be independent rv's with values in Ω ,
- $f: \Omega^n \rightarrow \mathbb{R}$ be \mathbf{c} -bounded on $\Xi \subset \Omega^n$, and
- $m = \mathbb{E} \{f(Z_1, \dots, Z_n) | (Z_1, \dots, Z_n) \in \Xi\}$, and
- $\gamma = 1 - \mathbb{P}\{(Z_1, \dots, Z_n) \in \Xi\}$.

Then for $\varepsilon > \gamma \|\mathbf{c}\|_1$

$$\mathbb{P} \{|f(Z_1, \dots, Z_n) - m| > \varepsilon\} \leq 2\gamma + 2 \exp \left(-\frac{2(\varepsilon - \gamma \|\mathbf{c}\|_1)^2}{\|\mathbf{c}\|_2^2} \right).$$



General framework

Define $\Xi = \Xi(C_1, C_2)$ as the set of $\mathbf{z} = (z_1, \dots, z_n)$ fulfilling

- 1 a uniform bound on the reconstruction error

$$\|R_h(\mathbf{z}_{-i}) - f\|_\infty \leq C_1$$

for $1 \leq i \leq n$.

General framework

Define $\Xi = \Xi(C_1, C_2)$ as the set of $\mathbf{z} = (z_1, \dots, z_n)$ fulfilling

- 1 a uniform bound on the reconstruction error

$$\|R_h(\mathbf{z}_{-i}) - f\|_\infty \leq C_1$$

- 2 $\mathbf{z}_{-i} \mapsto R_h(\mathbf{z}_{-i})(\mathbf{x})$ c -bounded for all $\mathbf{x} \in \Omega$ on Ξ with $c = \mathbb{1}C_2$

for $1 \leq i \leq n$.

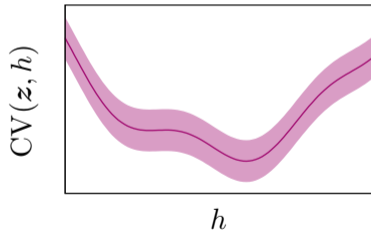
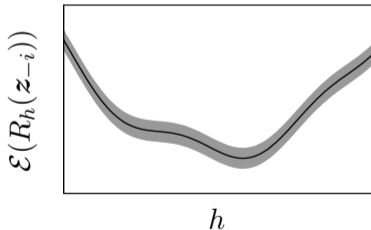
Risk functional $z \mapsto \mathcal{E}(R_h(z_{-i}))$ is c -bounded on Ξ with $c = 4C_1^2 \mathbf{1}$.

Cross-validation score $z \mapsto \text{CV}(z, h)$ is c -bounded on Ξ with $c = 5C_1^2 \mathbf{1}$.

Risk functional $z \mapsto \mathcal{E}(R_h(z_{-i}))$ is c -bounded on Ξ with $c = 4C_1^2 \mathbf{1}$.

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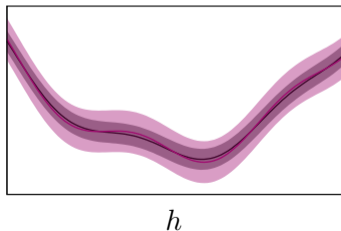
Now we apply Combes extension of McDiarmid.



Lemma

For \mathbf{Z}' representing $n - 1$ samples and \mathbf{Z} representing n samples element-wise distributed according to ρ we have

$$\mathbb{E}_{\mathbf{Z}'} \{ \mathcal{E}(R_h(\mathbf{Z}')) \} = \mathbb{E}_{\mathbf{Z}} \{ \text{CV}(\mathbf{Z}, h) \}.$$



$\mathcal{E}(R_h(\mathbf{z}_{-i}))$
 $\text{CV}(\mathbf{z}, h)$

Theorem (B., Hielscher 2021)

Let $n \geq 3$,

- \mathbf{Z} be element-wise distributed according to ρ with values in $(\Omega \times Y)^n$, and
- $\gamma = 1 - \mathbb{P}\{\mathbf{Z} \in \Xi\}$.

Then for $\varepsilon > 2\gamma \max\{5nC_1^2, (\|(R_h(\cdot))(\cdot)\|_\infty + \|f\|_\infty)^2\}$ we have

$$\begin{aligned} & \mathbb{P}\{|\text{CV}(\mathbf{Z}, h) - \mathcal{E}(R_h(\mathbf{Z}_{-1}))| > \varepsilon\} \\ & \leq 2\gamma + 2 \exp\left(-\left(\frac{\varepsilon}{12\sqrt{n}C_1^2} - \sqrt{2n\gamma}\right)^2\right). \end{aligned}$$

Shepard's model

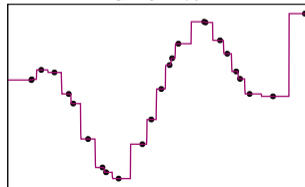
Shepard's model for $\Omega = \mathbb{T}$, $Y = \mathbb{R}$, and $\mathbf{z} = (x_i, f(x_i))_{i=1}^n$

$$R_h(\mathbf{z}) = \frac{\sum_{i=1}^m K_h(\cdot, x_i) f(x_i)}{\sum_{i=1}^m K_h(\cdot, x_i)}$$

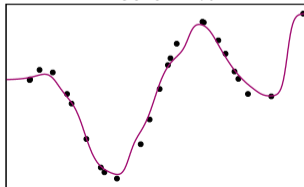
with

$$K_h(x_1, x_2) = \max\{0, 1 - h|x_1 - x_2|\}$$

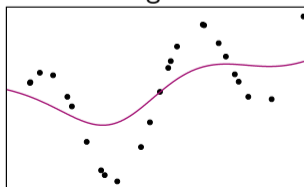
small h



medium h



big h



We need estimates for

- $\|(R_h(\cdot))(\cdot)\|_\infty$,
- $\|R_h(\mathbf{z}) - f\|_\infty$ for $\mathbf{z} \in \Xi$, and
- $\gamma = 1 - \mathbb{P}\{\mathbf{Z} \in \Xi\}$.

We need estimates for

- $\|(R_h(\cdot))(\cdot)\|_\infty \leq \|f\|_\infty$,
- $\|R_h(\mathbf{z}) - f\|_\infty$ for $\mathbf{z} \in \Xi$, and
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We need estimates for

- $\|(R_h(\cdot))(\cdot)\|_\infty \leq \|f\|_\infty$,
- $\|R_h(z) - f\|_\infty$ for $z \in \Xi$, and
- $\gamma = 1 - \mathbb{P}\{\mathbf{Z} \in \Xi\}$.

Theorem (B., Hielscher 2021)

Let x_1, \dots, x_n be uniformly random, $K_h(x, \cdot)$ be supported on $[x - 1/h, x + 1/h]$, and f be Lipschitz with constant L . Then $\|R_h(z) - f\|_\infty \leq L/h$ with probability

$$1 - \gamma \geq \sum_{k=0}^{\lfloor h \rfloor} (-1)^{k+1} \binom{n}{k} \left(1 - \frac{k}{2h}\right)^{n-1}$$

Theorem (B., Hielscher 2021)

Let

- f be Lipschitz with constant L ,
- $\mathbf{z} = (x_i, f(x_i))_{i=1}^n$ be samples with $x_1, \dots, x_n \in \mathbb{T}$ uniformly distributed,
- $K_h(x, \cdot)$ be supported on $[x - 1/h, x + 1/h]$, and
- $\gamma = \sum_{k=1}^{\lfloor h \rfloor} (-1)^{k+1} \binom{n}{k} \left(1 - \frac{k}{2h}\right)^{n-1}$.

Then we have for $\varepsilon > \gamma \max\{10nL^2/h^2, 8\|f\|_\infty^2\}$

$$\begin{aligned} & \mathbb{P} \{ |\text{CV}(\mathbf{Z}, h) - \mathcal{E}(R_h(\mathbf{Z}_{-1}))| > \varepsilon \} \\ & \leq 2\gamma + 2 \exp \left(- \left(\frac{h^2 \varepsilon}{12\sqrt{n}L^2} - \sqrt{2n\gamma} \right)^2 \right). \end{aligned}$$

Györfi, Kohler, Krzyżak, and Walk in 2002

Using

- binary kernels $K_h : \Omega \times \Omega \rightarrow \{0, 1\}$
- Ξ all possible samples

they showed

$$\gamma = 0, \quad C_1 = 2\|f\|_\infty, \quad C_2 \sim \frac{1}{n}$$

and obtained

$$\mathbb{P} \{ |\text{CV}(\mathbf{Z}, h) - \mathcal{E}(R_h(\mathbf{Z}_{-1}))| > \varepsilon \} \lesssim \exp \left(- (\sqrt{n}\varepsilon)^2 \right).$$

Györfi, Kohler, Krzyżak, and Walk in 2002

Using

- binary kernels $K_h : \Omega \times \Omega \rightarrow \{0, 1\}$
- Ξ all possible samples

they showed

$$\gamma = 0, \quad C_1 = 2\|f\|_\infty, \quad C_2 \sim \frac{1}{n}$$

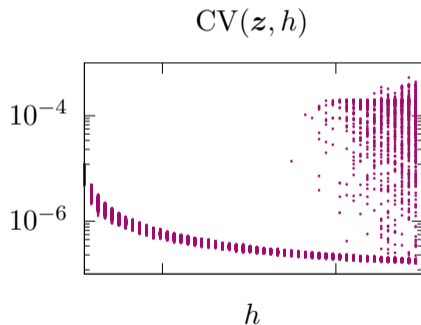
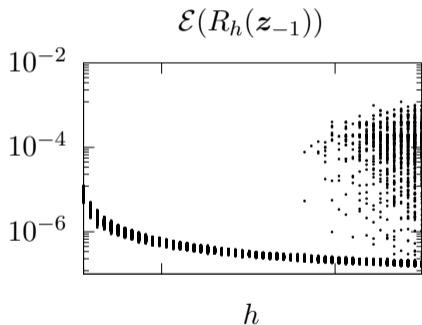
and obtained

$$\mathbb{P} \{ |\text{CV}(\mathbf{Z}, h) - \mathcal{E}(R_h(\mathbf{Z}_{-1}))| > \varepsilon \} \lesssim \exp \left(- (\sqrt{n}\varepsilon)^2 \right).$$

Our result assuming $h \sim n$:

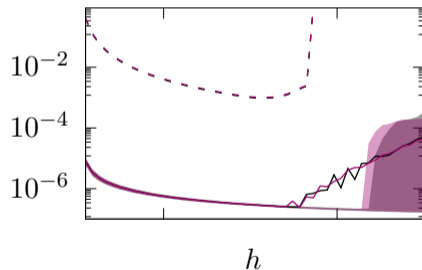
$$\mathbb{P} \{ |\text{CV}(\mathbf{Z}, h) - \mathcal{E}(R_h(\mathbf{Z}_{-1}))| > \varepsilon \} \lesssim \gamma + \exp \left(- \left(n^{3/2}\varepsilon + \sqrt{n}\gamma \right)^2 \right).$$

Numerical example



Numerical example

$CV(\mathbf{z}, h)$ and $\mathcal{E}(R_h(\mathbf{z}_{-1}))$



$|CV(\mathbf{z}, h) - \mathcal{E}(R_h(\mathbf{z}_{-1}))|$

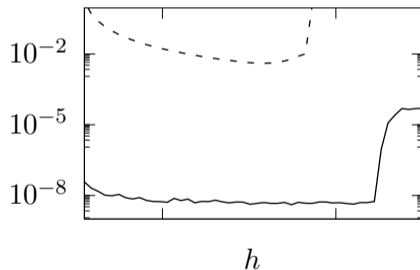


Table of contents

- ① Fast computation (for PLSE) ✓
- ② Theoretical foundation (for Shepard's model) ✓