

Optimality of Cross-validation in Scattered Data Approximation

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ESI Workshop on Tomographic Reconstructions and their Startling Applications

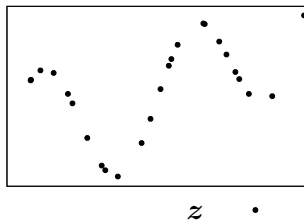
March 17, 2021



Mathematik!
TU Chemnitz

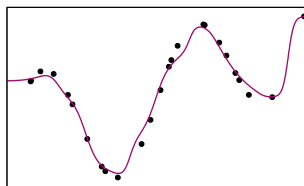
Setting

- data $z = ((x_i, y_i))_{i=1}^n \in (\Omega \times Y)^n$
- (x_i, y_i) distributed according to ρ



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- data $z = ((x_i, y_i))_{i=1}^n \in (\Omega \times Y)^n$
- (x_i, y_i) distributed according to ρ
- reconstruction algorithm
 $R_h: (\Omega \times Y)^n \rightarrow Y^\Omega$



$$R_h(z) \quad \bullet$$

Which reconstruction algorithm R_h to choose?

- risk functional

$$\mathcal{E}(R_h(\mathbf{z})) = \int_{\Omega \times Y} |(R_h(\mathbf{z}))(x) - y|^2 \, d\rho(x, y)$$

Which reconstruction algorithm R_h to choose?

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- cross-validation score (Golub, Heath, and Wahba '79)

$$\text{CV}(\mathbf{z}, h) = \frac{1}{n} \sum_{i=1}^n |(R_h(\mathbf{z}_{-i}))(x_i) - y_i|^2$$

Goal: bound their difference w.h.p.

Theorem (Hoeffding '63)

Let

- Z_1, \dots, Z_n independent rv's with values in $[0, 1]$ and
- $m = \mathbb{E} \left\{ \frac{1}{n} \sum_{i=1}^n Z_i \right\}$.

Then for $\varepsilon > 0$

$$\mathbb{P} \left\{ \left| \frac{1}{n} \sum_{i=1}^n Z_i - m \right| > \varepsilon \right\} \leq 2 \exp(-2n\varepsilon^2).$$

A function $f: \Omega^n \rightarrow \mathbb{R}$ is **c -bounded** on $\Xi \subset \Omega^n$ for $c = (c_1, \dots, c_n)$, iff

$$|f(z_1, \dots, z_n) - f(z_1, \dots, z_{i-1}, z'_i, z_{i+1}, \dots, z_n)| \leq c_i$$

for all $(z_1, \dots, z_n), (z'_1, \dots, z'_n) \in \Xi$, and $1 \leq i \leq n$.

Theorem (McDiarmid '89)

Let

- Z_1, \dots, Z_n be independent rv's with values in Ω ,
- $f: \Omega^n \rightarrow \mathbb{R}$ be c -bounded on Ω^n , and
- $m = \mathbb{E} \{f(Z_1, \dots, Z_n)\}$.

Then for $\varepsilon > 0$

$$\mathbb{P} \{|f(Z_1, \dots, Z_n) - m| > \varepsilon\} \leq 2 \exp\left(-\frac{2\varepsilon^2}{\|c\|_2^2}\right).$$

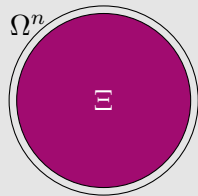
Theorem (Extension of McDiarmid by Richard Combes '15)

Let

- Z_1, \dots, Z_n be independent rv's with values in Ω ,
- $f: \Omega^n \rightarrow \mathbb{R}$ be \mathbf{c} -bounded on $\Xi \subset \Omega^n$, and
- $m = \mathbb{E} \{f(Z_1, \dots, Z_n) | (Z_1, \dots, Z_n) \in \Xi\}$, and
- $\gamma = 1 - \mathbb{P}\{(Z_1, \dots, Z_n) \in \Xi\}$.

Then for $\varepsilon > \gamma \|\mathbf{c}\|_1$

$$\mathbb{P} \{ |f(Z_1, \dots, Z_n) - m| > \varepsilon \} \leq 2\gamma + 2 \exp \left(-\frac{2(\varepsilon - \gamma \|\mathbf{c}\|_1)^2}{\|\mathbf{c}\|_2^2} \right).$$



Choice of Ξ

Define $\Xi = \Xi(C_1)$ as the set of $\mathbf{z} = (z_1, \dots, z_n)$ fulfilling

$$|R_h(\mathbf{z}_{-i})(x) - y| \leq C_1$$

for all $x \in \Omega, y \in Y$, and $1 \leq i \leq n$.

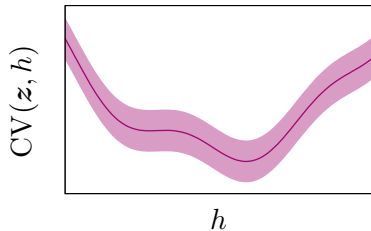
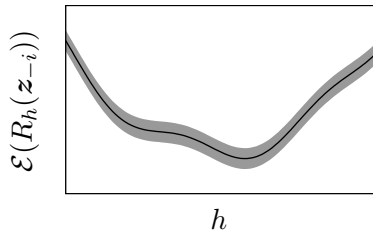
Risk functional $z \mapsto \mathcal{E}(R_h(z_{-i}))$ is c -bounded on Ξ with $c = 4C_1^2 \mathbf{1}$.

Cross-validation score $z \mapsto \text{CV}(z, h)$ is c -bounded on Ξ with $c = 5C_1^2 \mathbf{1}$.

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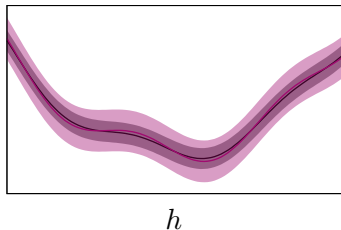
Now we apply Combes extension of McDiarmid.



Lemma

For \mathbf{Z}' representing $n - 1$ samples and \mathbf{Z} representing n samples element-wise distributed according to ρ , we have

$$\mathbb{E}_{\mathbf{Z}'} \{ \mathcal{E}(R_h(\mathbf{Z}')) \} = \mathbb{E}_{\mathbf{Z}} \{ \text{CV}(\mathbf{Z}, h) \}.$$



$\mathcal{E}(R_h(\mathbf{z}_{-i}))$
 $\text{CV}(\mathbf{z}, h)$

Theorem (B., Hielscher '21)

Let

- \mathbf{Z} be element-wise distributed according to ρ with values in $(\Omega \times Y)^n$ and
- $\gamma = 1 - \mathbb{P}\{\mathbf{Z} \in \Xi\}$.

Then for $\varepsilon > 2\gamma \max\{9nC_1^2, \|\text{CV}(\cdot, h)\|_\infty + \|\mathcal{E}(R_h(\cdot))\|_\infty\}$ and $n \geq 3$

$$\begin{aligned} & \mathbb{P}\{|\text{CV}(\mathbf{Z}, h) - \mathcal{E}(R_h(\mathbf{Z}_{-1}))| > \varepsilon\} \\ & \leq 2\gamma + 2 \exp\left(-\left(\frac{\varepsilon}{12\sqrt{n}C_1^2} - \sqrt{2n\gamma}\right)^2\right). \end{aligned}$$

Shepard's model

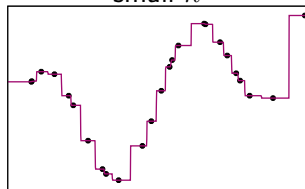
Shepard's model for $\Omega = \mathbb{T}$ and $Y = \mathbb{R}$ without noise

$$R_h(z) = \frac{\sum_{i=1}^m K_h(\cdot, x_i) f(x_i)}{\sum_{i=1}^m K_h(\cdot, x_i)}$$

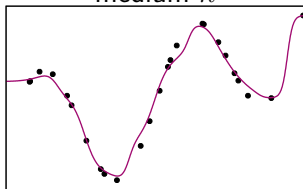
with

$$K_h(x_1, x_2) = \max\{0, 1 - h|x_1 - x_2|\}$$

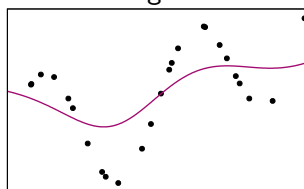
small h



medium h



big h



We need estimates for

- $\|CV(\cdot, h)\|_\infty$,
- $\|\mathcal{E}(R_h(\cdot))\|_\infty$,
- $C_1 = \max_{\mathbf{z} \in \Xi} \max_{x \in \Omega} |R_h(\mathbf{z})(x) - f(x)|$, and
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Theorem (B., Hielscher '21)

Let $K_h(x, \cdot)$ be supported on $[x - 1/h, x + 1/h]$ and f be Lipschitz with constant L . Then $C_1 > L/h$ with probability

$$\gamma = 1 - \left(1 - \left(1 - \frac{2}{h}\right)^{n-2} \frac{2n + h - 4}{h}\right)^n.$$

Theorem (B., Hielscher '21)

Let \mathbf{Z} represent n samples distributed according to ρ , $K_h(x, \cdot)$ be supported on $[x - 1/h, x + 1/h]$, and

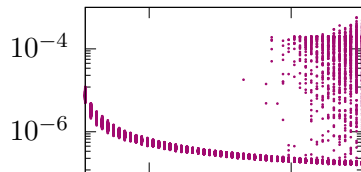
$$\gamma = 1 - \left(1 - \left(1 - \frac{2}{h} \right)^{n-2} \frac{2n + h - 4}{h} \right)^n.$$

We then have for $\varepsilon > \gamma \max\{10nL^2/h^2, 8\|f\|_\infty\}$

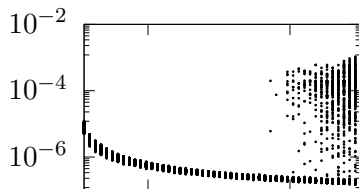
$$\begin{aligned} & \mathbb{P} \{ |\text{CV}(\mathbf{Z}, h) - \mathcal{E}(R_h(\mathbf{Z}_{-1}))| > \varepsilon \} \\ & \leq 2\gamma + 2 \exp \left(- \left(\frac{h^2 \varepsilon}{12\sqrt{n}L^2} - \sqrt{2n\gamma} \right)^2 \right). \end{aligned}$$

Numerical experiment for Shepard's model

$CV(z, h)$

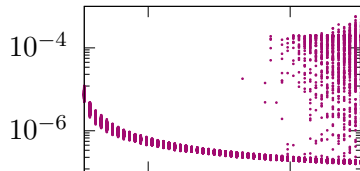


$\mathcal{E}(R_h(z_{-1}))$

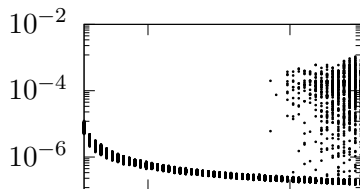


Numerical experiment for Shepard's model

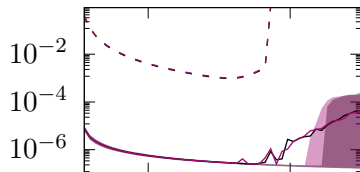
$CV(z, h)$



$\mathcal{E}(R_h(z_{-1}))$



$CV(z, h)$ and $\mathcal{E}(R_h(z_{-1}))$



$|CV(z, h) - \mathcal{E}(R_h(z_{-1}))|$

